

TESTING FOUR TREATMENTS USING 6 X 6 LATIN SQUARE

N a s r u l l a h *

R I N G K A S A N

Karangan ini menyetengahkan cara analisa statistik percobaan yang menguji empat perlakuan dengan rancangan lingkungan Rancangan Segi Empat Latin 6 x 6 yang setimbang baik pada baris maupun kolom. Rancangan ini dipakai dalam usaha memperoleh tambahan derajat bebas untuk mendapatkan nilai duga varians kesalahan percobaan yang memadai yang tidak mungkin dicapai dengan menduakalikan Rancangan Segi Empat Latin 4 x 4 tanpa mengorbankan pertimbangan lain yaitu ukuran terkecil petak yang diperlukan sehubungan dengan terbatasnya unit percobaan yang tersedia.

Suatu contoh numeris pemakaiannya juga diberikan.

S U M M A R Y

Statistical procedure for analyzing experiment testing four treatments in 6 x 6 Latin Square which is balanced in both row and column was described. The design is selected to get sufficient degrees of freedom to give an appropriate estimate of experimental error, since it is not possible to duplicate the 4 x 4 Latin Square without sacrificing another consideration, the minimum plot size, because of the limited experimental material.

A numerical example was given for clarification of the procedure.

I N T R O D U C T I O N

Trail and Weeks (1973) described a general statistical method for analyzing experiments arranged in an extended complete block design generated by balanced incomplete block design, a design in which some, but not all, treatments are duplicated within each block.

In field experiments, it is frequently happened that soil fertility and other variations in two directions are present. Likewise, in greenhouse experiments, there is variation along and across the greenhouse bench. If this is the case, Latin Square is an applicable skillful design.

Restriction of using Latin Square is on the requirement of having rows, columns, and treatments equally numerous to keep the design simple. Thus, to test four treatments, the square must be divided into four rows and four columns. The treatments are assigned so that each occurs once and only once in each row and each column. Due to few degrees of freedom available to give an appropriate estimate of experimental error, it is advised to double the 4 x 4 Latin Square (Kempthorne, 1960; Steel & Torrey, 1960). Sometimes it is happened that the square is somewhat large, in the sense that the division into rows and columns gives plot size larger than that required, but it is not possible to double the 4 x 4 Latin Square without sacrificing another consideration, namely the minimum plot size. The maximum number of rows and of columns the square can accommodate, taking the requirement of plot size into account is six. One method to solve the problem is introducing new treatments to make the number of treatments equals to the number of rows and of columns. But it is frequently not intended, as the new treatments are out of interest. The other method is using a balanced three way classification in which each row and each column contains one complete replicate of the treatments and a certain number of additional treatments, which are balanced in both rows and columns (Pearce, 1953).

Considering a statistical method for such experiments is the objective of the paper.

STATISTICAL ANALYSES

An example lay out to test four treatments using 6 x 6 Latin Square is illustrated in Fig. 1 (Pearce, 1953).

B	A	C	D	B	A
D	A	B	A	C	C
A	C	D	B	C	D
C	B	A	A	D	D
C	C	B	B	D	A
B	D	D	C	A	B

Fig. 1

In the model given below, ρ_i , γ_j , and δ_k are used to represent the effect of i^{th} row, j^{th} column, and k^{th} treatment respectively. Furthermore, it is assumed that there are no interaction among the three factors. Y_{ijk} is the yield of row i , column j which receives treatment k . Hence

$$Y = m j_1^{36} + X_1 \rho + X_2 \gamma + X_3 \delta + e \tag{1}$$

is the matrix model, where

$$\begin{matrix}
 \begin{bmatrix} Y_{112} \\ Y_{121} \\ \vdots \\ Y_{662} \end{bmatrix} & X_1 = & \begin{bmatrix} j_1^6 \\ \vdots \\ j_1^6 \end{bmatrix} & \rho = & \begin{bmatrix} \rho_1 \\ \rho_2 \\ \vdots \\ \rho_6 \end{bmatrix} & X_2 = & \begin{bmatrix} I_6 \\ \vdots \\ I_6 \end{bmatrix} \\
 \begin{bmatrix} \gamma_1 \\ \gamma_2 \\ \vdots \\ \gamma_6 \end{bmatrix} & X_3 = & \begin{bmatrix} 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 1 & 0 & 0 \end{bmatrix} & \delta = & \begin{bmatrix} \delta_1 \\ \delta_2 \\ \vdots \\ \delta_4 \end{bmatrix} & e = & \begin{bmatrix} e_{112} \\ e_{121} \\ \vdots \\ e_{662} \end{bmatrix}
 \end{matrix}$$

The constants associated with the design are : number of rows, number of columns, number of treatments, number of application of each treatment, number of distinct pairs of experimental units which receive any fixed pair of treatments while appearing in the same row and that in the same column. They are denoted by r, c, t, s, λ_r and λ_c which are equal to 6, 6, 4, 9, 13, and 13 respectively.

The error sum of squares is

$$e'e = Y'^* Y^* - 2Y'^* X_1 \rho - 2Y'^* X_2 \gamma - 2Y'^* X_3 \zeta + \rho' X_1' X_1 \rho + \gamma' X_2' X_2 \gamma + \zeta' X_3' X_3 \zeta + 2\rho' X_1' X_2 \gamma + 2\rho' X_1' X_3 \zeta + 2\gamma' X_2' X_3 \zeta \quad (2)$$

where

$$Y^* = Y - mj_1 \quad (3)$$

Differentiating (2) with respect to ρ , γ , and ζ alternatively, and each is equated to zero, we get after dividing by 2

$$X_1' Y^* = X_1' X_1 \rho + X_1' X_2 \gamma + X_1' X_3 \zeta \quad (4)$$

$$X_2' Y^* = X_2' X_1 \rho + X_2' X_2 \gamma + X_2' X_3 \zeta \quad (5)$$

$$X_3' Y^* = X_3' X_1 \rho + X_3' X_2 \gamma + X_3' X_3 \zeta \quad (6)$$

Let $\hat{\rho}$, $\hat{\gamma}$, and $\hat{\zeta}$ is the estimate of ρ , γ , and ζ . By the use of Gauss sweep out method, the normal equations under restrictions of $j'\hat{\rho} = 0$, $j'\hat{\gamma} = 0$, and $j'\hat{\zeta} = 0$ become

$$\hat{\rho} = (X_1' X_1)^{-1} (X_1' Y^* - X_3' \hat{\zeta}) \quad (7)$$

$$\hat{\gamma} = (X_2' X_2)^{-1} (X_2' Y^* - X_3' \hat{\zeta}) \quad (8)$$

$$\hat{\zeta} = [X_3' X_3 - X_3' X_1 (X_1' X_1)^{-1} X_1' X_3 - X_3' X_2 (X_2' X_2)^{-1} X_2' X_3]^{-1} [X_3' Y^* - X_3' X_1 (X_1' X_1)^{-1} X_1' Y^* - X_3' X_2 (X_2' X_2)^{-1} X_2' Y^*] \quad (9)$$

Substituting (4), (5), and (6) to (2) gives

$$e'e = Y'^* Y^* - Y'^* X_1 \hat{\rho} - Y'^* X_2 \hat{\gamma} - Y'^* X_3 \hat{\zeta} \quad (10)$$

and using (7) and (8), (10) becomes

$$e'e = Y'^* Y^* - Y'^* X_1 (X_1' X_1)^{-1} X_1' Y^* - Y'^* X_2 (X_2' X_2)^{-1} X_2' Y^* - [Y'^* X_3 - Y'^* X_1 (X_1' X_1)^{-1} X_1' X_3 - Y'^* X_2 (X_2' X_2)^{-1} X_2' X_3] \hat{\zeta}$$

Finally, inserting (9) we get

$$e'e = Y'^* Y^* - Y'^* X_1 (X_1' X_1)^{-1} X_1' Y^* - Y'^* X_2 (X_2' X_2)^{-1} X_2' Y^* - [X_3' Y^* - X_3' X_1 (X_1' X_1)^{-1} X_1' Y^* - X_3' X_2 (X_2' X_2)^{-1} X_2' Y^*] [X_3' X_3 - X_3' X_1 (X_1' X_1)^{-1} X_1' X_3 - X_3' X_2 (X_2' X_2)^{-1} X_2' X_3]^{-1} [X_3' Y^* - X_3' X_1 (X_1' X_1)^{-1} X_1' Y^* - X_3' X_2 (X_2' X_2)^{-1} X_2' Y^*] \quad (11)$$

Denoting $A = I - X_1 (X_1' X_1)^{-1} X_1' - X_2 (X_2' X_2)^{-1} X_2'$, the analysis of variance is shown in Table 1.

Table 1. Analysis of Variance

Source of Variation	df	Sum of Square	E (MS)
Row (unadj.)	$r - 1$	$Y^* ' X_1 (X_1' X_1)^{-1} X_1' Y^*$	
Column (unadj.)	$r - 1$	$Y^* ' X_2 (X_2' X_2)^{-1} X_2' Y^*$	
Treatment (adj.)	$t - 1$	$Y^* ' A X_3 (X_3' A X_3)^{-1} X_3' A Y^*$	$\sigma_e^2 + \frac{2\lambda t - sr}{r(t-1)} \sum (\bar{y}_k - \bar{y})^2$
Error	$(r-1)^2$ $-(t-1)$	By subtraction	σ_e^2
Total	$r^2 - 1$	$Y^* ' Y^*$	

ESTIMATION OF TREATMENT EFFECTS

Recall equation (9), $C \hat{\bar{y}} = X_3' Y^* - X_3' X_1 (X_1' X_1)^{-1} X_1' Y^* - X_3' X_2 (X_2' X_2)^{-1} X_2' Y^*$ where $C = X_3' X_3 - X_3' X_1 (X_1' X_1)^{-1} X_1' X_2 - X_3' X_2 (X_2' X_2)^{-1} X_2' X_3$. As consequence of the earlier definitions, it may be seen that

- (9) $X_1' X_1 = rI$
- $X_2' X_2 = cI$
- $X_3' X_3 = tI$
- (10) $X_1' X_2 = J$
- $X_1' X_3 = N$ (incidence matrix ignoring the column)
- $X_2' X_3 = M$ (incidence matrix ignoring the row)
- $r = c$
- $rc = st$
- $\lambda_c = \lambda_r = \lambda$
- $N'N = J + (rs - t) I$
- $M'M = J + (cs - t) I$
- $N'N = M'M$

So
$$C = \frac{2\lambda t - sr}{r} \left[I - \frac{2}{2\lambda t - sr} J \right]$$

and the variance of the difference of the two treatment means is

$$V \text{ ar } (\hat{\bar{y}}_k - \hat{\bar{y}}_{k'}) = \frac{2r}{2\lambda t - sr} \sigma_e^2$$

NUMERICAL EXAMPLE

The table below is a hypothetical yield of experiment whose lay out is shown in Fig. 1.

B = 55.8	A = 28.0	C = 45.3	D = 47.2	B = 52.4	A = 29.6
D = 47.2	A = 28.8	B = 48.5	A = 34.8	C = 47.4	C = 48.2
A = 33.7	C = 43.5	D = 46.3	B = 48.4	C = 39.1	D = 45.5
C = 41.0	B = 40.8	A = 41.7	A = 42.0	D = 41.7	D = 42.5
C = 38.5	C = 38.2	B = 36.7	B = 35.2	D = 41.6	A = 37.0
B = 41.2	D = 33.9	D = 35.3	C = 36.2	A = 45.2	B = 39.4

Total of rows (R_i), total of column (C_j), total of treatment (T_k), and general total are

$R_1 = 258.3$	$C_1 = 257.4$	$T_1 = 320.8$	$GT = 1,477.8$
$R_2 = 254.9$	$C_2 = 213.2$	$T_2 = 398.4$	
$R_3 = 256.5$	$C_3 = 253.8$	$T_3 = 377.4$	
$R_4 = 249.7$	$C_4 = 243.8$	$T_4 = 381.4$	
$R_5 = 227.2$	$C_5 = 267.4$		
$R_6 = 231.2$	$C_6 = 242.2$		
$m = \frac{1,477.8}{36}$			

$$CF = \frac{(1,477.8)^2}{36} = 60,663.69$$

Total Sum of Squares

$$= (55.8)^2 + (47.2)^2 + \dots + (39.4)^2 - CF$$

$$= 62,122.54 - 60,663.69$$

$$R_{SS} = \frac{(258.3)^2 + (254.9)^2 + \dots + (231.2)^2}{6} - CF$$

$$= 60,818.09 - 60,663.69$$

$$C_{SS} = \frac{(257.4)^2 + (213.2)^2 + \dots + (242.2)^2}{6} - CF$$

$$= 60,954.25 - 60,663.69$$

$$= 290.56$$

$X_3' Y^* - X_3' X_1 (X_1' X_1)^{-1} X_1' Y^* - X_3' X_2 (X_2' X_2)^{-1} X_2' Y^*$ is

$$\begin{bmatrix} 320.8 \\ 398.4 \\ 377.4 \\ 381.2 \end{bmatrix} - \frac{1}{6} \begin{bmatrix} 2 & 2 & 1 & 2 & 1 & 1 \\ 2 & 1 & 1 & 1 & 2 & 2 \\ 1 & 2 & 2 & 1 & 2 & 1 \\ 1 & 1 & 2 & 2 & 1 & 2 \end{bmatrix} \begin{bmatrix} 258.3 \\ 254.9 \\ 256.5 \\ 249.7 \\ 227.2 \\ 231.2 \end{bmatrix} - \frac{1}{6} \begin{bmatrix} 1 & 2 & 1 & 2 & 1 & 2 \\ 2 & 1 & 2 & 2 & 1 & 1 \\ 2 & 2 & 1 & 1 & 2 & 1 \\ 1 & 1 & 2 & 1 & 2 & 2 \end{bmatrix} \begin{bmatrix} 257.4 \\ 213.2 \\ 253.8 \\ 243.8 \\ 267.4 \\ 242.2 \end{bmatrix} + \begin{bmatrix} 369.45 \\ 369.45 \\ 369.45 \\ 369.45 \end{bmatrix}$$

and equal to

$$T_{SS} = \frac{6}{50} \begin{bmatrix} -46.03 \\ 29.97 \\ 8.15 \\ 7.92 \end{bmatrix} \begin{bmatrix} -46.03 \\ 29.97 \\ 8.15 \\ 7.92 \end{bmatrix} = 377.54$$

$$E_{SS} = 1,458.85 - 154.40 - 290.56 - 377.54 \\ = 636.35$$

The Analysis of Variance is shown below

Analysis of Variance				
Source of Var.	df	SS	MS	F
Row (unadj.)	5	154.40		
Column (unadj.)	5	290.56		
Treatment (adj.)	3	377.54	125.85	4.35**
Error	22	636.35	28.93	
Total	35	1,458.85		

** Significant at 1% level of significance

Estimation of the treatment effects is shown below

$$\hat{\tau} = \frac{6}{50} \begin{bmatrix} -46.03 \\ 29.97 \\ 8.15 \\ 7.92 \end{bmatrix} = \begin{bmatrix} -5.524 \\ 3.596 \\ .978 \\ .950 \end{bmatrix}$$

with variance of two treatments difference

$$\text{Var} (\hat{\tau}_k - \hat{\tau}_{k'}) = \frac{12}{50} (28.93) \\ = 6.94$$

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